

**CIMPA SCHOOL, 2007**  
**Jump Processes and Applications to Finance**  
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## **Jump Processes**

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## **I. Poisson Processes**

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# 1 Counting Processes and Stochastic Integral

Let  $T_n$  be an increasing sequence of random times.

$$N_t = \begin{cases} n & \text{if } t \in [T_n, T_{n+1}[ \\ +\infty & \text{otherwise} \end{cases}$$

or, equivalently

$$N_t = \sum_{n \geq 1} \mathbb{1}_{\{T_n \leq t\}} = \sum_{n \geq 1} n \mathbb{1}_{\{T_n \leq t < T_{n+1}\}}.$$

We denote by  $N_{t-}$  the left-limit of  $N_s$  when  $s \rightarrow t, s < t$  and by  $\Delta N_s = N_s - N_{s-}$  the jump process of  $N$ .

The stochastic integral

$$\int_0^t C_s dN_s$$

is defined as

$$(C \star N)_t = \int_0^t C_s dN_s = \int_{]0,t]} C_s dN_s = \sum_{n=1}^{\infty} C_{T_n} \mathbb{1}_{\{T_n \leq t\}}.$$

## 2 Standard Poisson Process

### 2.1 Definition

The standard Poisson process is a counting process such that

- for every  $s, t$ ,  $N_{t+s} - N_t$  is independent of  $\mathcal{F}_t^N$ ,
- for every  $s, t$ , the r.v.  $N_{t+s} - N_t$  has the same law as  $N_s$ .

Then,

$$P(N_t = k) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

## 2.2 First properties

- $E(N_t) = \lambda t, \quad \text{Var}(N_t) = \lambda t$
- for every  $x > 0, t > 0, u, \alpha \in \mathbb{R}$

$$E(x^{N_t}) = e^{\lambda t(x-1)}; \quad E(e^{iuN_t}) = e^{\lambda t(e^{iu}-1)}; \quad E(e^{\alpha N_t}) = e^{\lambda t(e^{\alpha}-1)}.$$

## 2.3 Martingale properties

For each  $\alpha \in \mathbb{R}$ , for each bounded Borel function  $h$ , the following processes are  $\mathbf{F}$ -martingales:

$$(i) \quad M_t = N_t - \lambda t,$$

$$(ii) \quad M_t^2 - \lambda t = (N_t - \lambda t)^2 - \lambda t,$$

$$(iii) \quad \exp(\alpha N_t - \lambda t(e^\alpha - 1)),$$

$$(iv) \quad \exp\left[\int_0^t h(s)dN_s - \lambda \int_0^t (e^{h(s)} - 1)ds\right].$$

For any  $\beta > -1$ , any bounded Borel function  $h$ , and any bounded Borel function  $\varphi$  valued in  $] - 1, \infty[$ , the processes

$$\begin{aligned} \exp[\ln(1 + \beta)N_t - \lambda\beta t] &= (1 + \beta)^{N_t} e^{-\lambda\beta t}, \\ \exp\left[\int_0^t h(s)dM_s + \lambda \int_0^t (1 + h(s) - e^{h(s)})ds\right], \\ \exp\left[\int_0^t \ln(1 + \varphi(s))dN_s - \lambda \int_0^t \varphi(s)ds\right], \\ \exp\left[\int_0^t \ln(1 + \varphi(s))dM_s + \lambda \int_0^t (\ln(1 + \varphi(s)) - \varphi(s)ds)\right], \end{aligned}$$

are martingales.

Let  $H$  be an  $\mathbf{F}$ -predictable bounded process, then the following processes are martingales

$$\begin{aligned} (H \star M)_t &:= \int_0^t H_s dM_s = \int_0^t H_s dN_s - \lambda \int_0^t H_s ds \\ ((H \star M)_t)^2 - \lambda \int_0^t H_s^2 ds \\ \exp \left( \int_0^t H_s dN_s + \lambda \int_0^t (1 - e^{H_s}) ds \right) \end{aligned}$$

These properties does not extend to adapted processes  $H$ . For example, from  $\int_0^t (N_s - N_{s-}) dM_s = N_t$ , it follows that  $\int_0^t N_s dM_s$  is not a martingale.

**Remark 2.1** Note that (i) and (iii) imply that the process  $(M_t^2 - N_t; t \geq 0)$  is a martingale.

The process  $\lambda t$  is the predictable quadratic variation  $\langle M \rangle$ , whereas the process  $(N_t, t \geq 0)$  is the optional quadratic variation  $[M]$ .

For any  $\mu \in [0, 1]$ , the processes  $M_t^2 - (\mu N_t + (1 - \mu)\lambda t)$  are also martingales.

## 2.4 Infinitesimal Generator

The Poisson process is a Lévy process, hence a Markov process, its infinitesimal generator  $\mathcal{L}$  is defined as

$$\mathcal{L}(f)(x) = \lambda[f(x+1) - f(x)].$$

Therefore, for any bounded Borel function  $f$ , the process

$$C_t^f = f(N_t) - f(0) - \int_0^t \mathcal{L}(f)(N_s) ds$$

is a martingale.

Furthermore,

$$C_t^f = \int_0^t [f(N_{s-} + 1) - f(N_{s-})] dM_s.$$

Exercise:

Extend the previous formula to functions  $f(t, x)$  defined on  $\mathbb{R}^+ \times \mathbb{R}$  and  $C^1$  with respect to  $t$ , and prove that if

$$L_t = \exp(\log(1 + \phi)N_t - \lambda\phi t)$$

then

$$dL_t = L_{t-} \phi dM_t .$$

### 2.4.1 Watanabe's characterization

Let  $N$  be a counting process and assume that there exists  $\lambda > 0$  such that  $M_t = N_t - \lambda t$  is a martingale. Then  $N$  is a Poisson process with intensity  $\lambda$ .

## 2.5 Change of Probability

If  $N$  is a Poisson process, then, for  $\beta > -1$ ,

$$L_t = (1 + \beta)^{N_t} e^{-\lambda\beta t}$$

is a strictly positive martingale with expectation equal to 1.

Let  $Q$  be the probability defined as  $\frac{dQ}{dP}|_{\mathcal{F}_t} = L_t$ .

The process  $N$  is a  $Q$ -Poisson process with intensity equal to  $(1 + \beta)\lambda$ .

## 2.6 Hitting Times

Let  $T_x = \inf\{t, N_t \geq x\}$ . Then, for  $n \leq x < n + 1$ ,  $T_x$  is equal to the time of the  $n^{\text{th}}$ -jump of  $N$ , hence has a Gamma ( $n$ ) law.

# 3 Inhomogeneous Poisson Processes

## 3.1 Definition

Let  $\lambda$  be an  $\mathbb{R}^+$ -valued function satisfying  $\int_0^t \lambda(u) du < \infty, \forall t$ .

An inhomogeneous Poisson process  $N$  with intensity  $\lambda$  is a counting process with independent increments which satisfies for  $t > s$

$$P(N_t - N_s = n) = e^{-\Lambda(s,t)} \frac{(\Lambda(s,t))^n}{n!}$$

where  $\Lambda(s,t) = \Lambda(t) - \Lambda(s) = \int_s^t \lambda(u) du$ , and  $\Lambda(t) = \int_0^t \lambda(u) du$ .

For any real numbers  $u$  and  $\alpha$ , for any  $t$

$$E(e^{iuN_t}) = \exp((e^{iu} - 1)\Lambda(t))$$

$$E(e^{\alpha N_t}) = \exp((e^{\alpha} - 1)\Lambda(t)).$$

## 3.2 Martingale Properties

Let  $N$  be an inhomogeneous Poisson process with deterministic intensity  $\lambda$ . The process

$$(M_t = N_t - \int_0^t \lambda(s)ds, t \geq 0)$$

is an  $\mathbf{F}^N$ -martingale, and the increasing function  $\Lambda(t) = \int_0^t \lambda(s)ds$  is called the compensator of  $N$ .

Let  $\phi$  be an  $\mathbf{F}^N$ -predictable process such that  $E(\int_0^t |\phi_s| \lambda(s)ds) < \infty$  for every  $t$ . Then, the process  $(\int_0^t \phi_s dM_s, t \geq 0)$  is an  $\mathbf{F}^N$ -martingale.

In particular,

$$E \left( \int_0^t \phi_s dN_s \right) = E \left( \int_0^t \phi_s \lambda(s) ds \right) .$$

Let  $H$  be an  $\mathbf{F}^N$ -predictable process. The following processes are martingales

$$\begin{aligned} (H \star M)_t &= \int_0^t H_s dM_s = \int_0^t H_s dN_s - \int_0^t \lambda(s) H_s ds \\ ((H \star M)_t)^2 &- \int_0^t \lambda(s) H_s^2 ds \\ \exp \left( \int_0^t H_s dN_s - \int_0^t \lambda(s) (e^{H_s} - 1) ds \right) &. \end{aligned}$$

### 3.3 Watanabe's Characterization

**Proposition 3.1 (Watanabe characterization.)** *Let  $N$  be a counting process and  $\Lambda$  an increasing, continuous function with zero value at time zero. Let us assume that  $M_t = N_t - \Lambda(t)$  is a martingale. Then  $N$  is an inhomogeneous Poisson process with compensator  $\Lambda$ .*

## 3.4 Stochastic calculus

### 3.4.1 Integration by parts formula

Let  $X_t = x + \int_0^t x_s dN_s$  and  $Y_t = y + \int_0^t y_s dN_s$ , where  $x$  and  $y$  are predictable processes.

$$\begin{aligned} X_t Y_t &= xy + \sum_{s \leq t} \Delta(XY)_s = xy + \sum_{s \leq t} Y_{s-} \Delta X_s + \sum_{s \leq t} X_{s-} \Delta Y_s + \sum_{s \leq t} \Delta X_s \Delta Y_s \\ &= xy + \int_0^t X_{s-} dY_s + \int_0^t Y_{s-} dX_s + [X, Y]_t \end{aligned}$$

where

$$[X, Y]_t = \sum_{s \leq t} \Delta X_s \Delta Y_s = \sum_{s \leq t} x_s y_s \Delta N_s = \int_0^t x_s y_s dN_s.$$

If  $dX_t = \mu_t dt + x_t dN_t$  and  $dY_t = \nu_t dt + y_t dN_t$ , one gets

$$X_t Y_t = xy + \int_0^t Y_{s-} dX_s + \int_0^t Y_{s-} dX_s + [X, Y]_t$$

where

$$[X, Y]_t = \int_0^t x_s y_s dN_s .$$

If  $dX_t = x_t dM_t$  and  $dY_t = y_t dM_t$ , the process  $X_t Y_t - [X, Y]_t$  is a martingale.

### 3.4.2 Itô's Formula

Let  $N$  be a Poisson process and  $f$  a bounded Borel function. The decomposition

$$f(N_t) = f(N_0) + \sum_{0 < s \leq t} [f(N_s) - f(N_{s-})]$$

is trivial and corresponds to Itô's formula for a Poisson process.

$$\begin{aligned} \sum_{0 < s \leq t} [f(N_s) - f(N_{s-})] &= \sum_{0 < s \leq t} [f(N_{s-} + 1) - f(N_{s-})] \Delta N_s \\ &= \int_0^t [f(N_{s-} + 1) - f(N_{s-})] dN_s . \end{aligned}$$

Let

$$X_t = x + \int_0^t x_s dN_s = x + \sum_{T_n \leq t} x_{T_n} ,$$

with  $x$  a predictable process.

The trivial equality

$$F(X_t) = F(X_0) + \sum_{s \leq t} (F(X_s) - F(X_{s-})) ,$$

holds for any bounded function  $F$ .

Let

$$dX_t = \mu_t dt + x_t dM_t = (\mu_t - x_t \lambda(t)) dt + x_t dN_t$$

and  $F \in C^{1,1}(\mathbb{R}^+ \times \mathbb{R})$ . Then

$$\begin{aligned}
F(t, X_t) &= F(0, X_0) + \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) (\mu_s - x_s \lambda(s)) ds \\
&\quad + \sum_{s \leq t} F(s, X_s) - F(s, X_{s-}) \tag{3.1} \\
&= F(0, X_0) + \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\
&\quad + \sum_{s \leq t} [F(s, X_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) x_s \Delta N_s] .
\end{aligned}$$

The formula (3.1) can be written as

$$\begin{aligned}
F(t, X_t) - F(0, X_0) &= \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_s) (\mu_s - x_s \lambda(s)) ds \\
&\quad + \int_0^t [F(s, X_s) - F(s, X_{s-})] dN_s \\
&= \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\
&\quad + \int_0^t [F(s, X_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) x_s] dN_s \\
&= \int_0^t \partial_t F(s, X_s) ds + \int_0^t \partial_x F(s, X_{s-}) dX_s \\
&\quad + \int_0^t [F(s, X_{s-} + x_s) - F(s, X_{s-}) - \partial_x F(s, X_{s-}) x_s] dN_s .
\end{aligned}$$

## 3.5 Predictable Representation Property

**Proposition 3.2** *Let  $\mathbf{F}^N$  be the completion of the canonical filtration of the Poisson process  $N$  and  $H \in L^2(\mathcal{F}_\infty^N)$ , a square integrable random variable. Then, there exists a predictable process  $h$  such that*

$$H = E(H) + \int_0^\infty h_s dM_s$$

*and  $E(\int_0^\infty h_s^2 ds) < \infty$ .*

## 3.6 Independent Poisson Processes

**Definition 3.3** A process  $(N^1, \dots, N^d)$  is a  $d$ -dimensional  $\mathbf{F}$ -Poisson process if each  $N^j$  is a right-continuous adapted process such that  $N_0^j = 0$  and if there exists constants  $\lambda_j$  such that for every  $t \geq s \geq 0$

$$P \left[ \bigcap_{j=1}^d (N_t^j - N_s^j = n_j) \mid \mathcal{F}_s \right] = \prod_{j=1}^d e^{-\lambda_j(t-s)} \frac{(\lambda_j(t-s))^{n_j}}{n_j!}.$$

**Proposition 3.4** An  $\mathbf{F}$ -adapted process  $N$  is a  $d$ -dimensional  $\mathbf{F}$ -Poisson process if and only if

- (i) each  $N^j$  is an  $\mathbf{F}$ -Poisson process
- (ii) no two  $N^j$ 's jump simultaneously.

# 4 Stochastic intensity processes

## 4.1 Definition

**Definition 4.1** *Let  $N$  be a counting process,  $\mathbf{F}$ -adapted and  $(\lambda_t, t \geq 0)$  a non-negative  $\mathbf{F}$ - progressively measurable process such that for every  $t$ ,  $\Lambda_t = \int_0^t \lambda_s ds < \infty$   $P$  a.s..*

*The process  $N$  is an inhomogeneous Poisson process with stochastic intensity  $\lambda$  if for every non-negative  $\mathbf{F}$ -predictable process  $(\phi_t, t \geq 0)$  the following equality is satisfied*

$$E \left( \int_0^\infty \phi_s dN_s \right) = E \left( \int_0^\infty \phi_s \lambda_s ds \right) .$$

Therefore  $(M_t = N_t - \Lambda_t, t \geq 0)$  is an  $\mathbf{F}$ -local martingale.

If  $\phi$  is a predictable process such that  $\forall t, E(\int_0^t |\phi_s| \lambda_s ds) < \infty$ , then  $(\int_0^t \phi_s dM_s, t \geq 0)$  is an  $\mathbf{F}$ -martingale.

An inhomogeneous Poisson process  $N$  with stochastic intensity  $\lambda_t$  can be viewed as time changed of  $\tilde{N}$ , a standard Poisson process  $N_t = \tilde{N}_{\Lambda_t}$ .

## 4.2 Itô's formula

The formula obtained in Section 3.4 generalizes to inhomogeneous Poisson process with stochastic intensity.

### 4.3 Exponential Martingales

**Proposition 4.2** *Let  $N$  be an inhomogeneous Poisson process with stochastic intensity  $(\lambda_t, t \geq 0)$ , and  $(\mu_t, t \geq 0)$  a predictable process such that  $\int_0^t |\mu_s| \lambda_s ds < \infty$ . Then, the process  $L$  defined by*

$$L_t = \begin{cases} \exp(-\int_0^t \mu_s \lambda_s ds) & \text{if } t < T_1 \\ \prod_{n, T_n \leq t} (1 + \mu_{T_n}) \exp(-\int_0^t \mu_s \lambda_s ds) & \text{if } t \geq T_1 \end{cases} \quad (4.1)$$

*is a local martingale, solution of*

$$dL_t = L_{t-} \mu_t dM_t, \quad L_0 = 1. \quad (4.2)$$

*Moreover, if  $\mu$  is such that  $\forall s, \mu_s > -1$ ,*

$$L_t = \exp \left[ -\int_0^t \mu_s \lambda_s ds + \int_0^t \ln(1 + \mu_s) dN_s \right].$$

The local martingale  $L$  is denoted by  $\mathcal{E}(\mu \star M)$  and named the Doléans-Dade exponential of the process  $\mu \star M$ .

This process can also be written

$$L_t = \prod_{0 < s \leq t} (1 + \mu_s \Delta N_s) \exp \left[ - \int_0^t \mu_s \lambda_s ds \right].$$

Moreover, if  $\forall t, \mu_t > -1$ , then  $L$  is a non-negative local martingale, therefore it is a supermartingale and

$$\begin{aligned} L_t &= \exp \left[ - \int_0^t \mu_s \lambda_s ds + \sum_{s \leq t} \ln(1 + \mu_s) \Delta N_s \right] \\ &= \exp \left[ - \int_0^t \mu_s \lambda_s ds + \int_0^t \ln(1 + \mu_s) dN_s \right] \\ &= \exp \left[ \int_0^t [\ln(1 + \mu_s) - \mu_s] \lambda_s ds + \int_0^t \ln(1 + \mu_s) dM_s \right]. \end{aligned}$$

The process  $L$  is a martingale if  $\forall t, E(L_t) = 1$ .

If  $\mu$  is not greater than  $-1$ , then the process  $L$  defined in (4.1) is still a local martingale which satisfies  $dL_t = L_{t-}\mu_t dM_t$ . However it may be negative.

## 4.4 Change of Probability

Let  $\mu$  be a predictable process such that  $\mu > -1$  and  $\int_0^t \lambda_s |\mu_s| ds < \infty$ .

Let  $L$  be the positive exponential local martingale solution of

$$dL_t = L_{t-} \mu_t dM_t .$$

Assume that  $L$  is a martingale and let  $Q$  be the probability measure equivalent to  $P$  defined on  $\mathcal{F}_t$  by  $Q|_{\mathcal{F}_t} = L_t P|_{\mathcal{F}_t}$ .

Under  $Q$ , the process

$$(M_t^\mu \stackrel{def}{=} M_t - \int_0^t \mu_s \lambda_s ds = N_t - \int_0^t (\mu_s + 1) \lambda_s ds , t \geq 0)$$

is a local martingale.

# 5 Compound Poisson Processes

## 5.1 Definition and Properties

Let  $\lambda$  be a positive number and  $F(dy)$  be a probability law on  $\mathbb{R}$ . A  $(\lambda, F)$  compound Poisson process is a process  $X = (X_t, t \geq 0)$  of the form

$$X_t = \sum_{k=1}^{N_t} Y_k$$

where  $N$  is a Poisson process with intensity  $\lambda > 0$  and the  $(Y_k, k \in \mathbb{N})$  are i.i.d. square integrable random variables with law  $F(dy) = P(Y_1 \in dy)$ , independent of  $N$ .

**Proposition 5.1** *A compound Poisson process has stationary and independent increments; the cumulative function of  $X_t$  is*

$$P(X_t \leq x) = e^{-\lambda t} \sum_n \frac{(\lambda t)^n}{n!} F^{*n}(x).$$

*If  $E(|Y_1|) < \infty$ , the process  $(Z_t = X_t - t\lambda E(Y_1), t \geq 0)$  is a martingale and in particular,  $E(X_t) = \lambda t E(Y_1) = \lambda t \int y F(dy)$ .*

*If  $E(Y_1^2) < \infty$ , the process  $(Z_t^2 - t\lambda E(Y_1^2), t \geq 0)$  is a martingale and  $\text{Var}(X_t) = \lambda t E(Y_1^2)$ .*

Introducing the random measure  $\mu = \sum_n \delta_{T_n, Y_n}$  on  $\mathbb{R}^+ \times \mathbb{R}$  and denoting by  $(f * \mu)_t$  the integral  $\int_0^t \int_{\mathbb{R}} f(x) \mu(\omega, ds, dx)$ , we obtain that

$$M_t^f = (f * \mu)_t - t\nu(f) = \int_0^t \int_{\mathbb{R}} f(x) (\mu(\omega, ds, dx) - ds\nu(dx))$$

is a martingale.

**Corollary 5.2** *Let  $X$  be a  $(\lambda, F)$  compound Poisson process independent of  $W$ . Let*

$$dS_t = S_{t-}(\mu dt + dX_t).$$

*Then,*

$$S_t = S_0 e^{\mu t} \prod_{k=1}^{N_t} (1 + Y_k)$$

*In particular, if  $1 + Y_1 > 0$ , a.s.*

$$S_t = S_0 \exp\left(\mu t + \sum_{k=1}^{N_t} \ln(1 + Y_k)\right).$$

*The process  $(S_t e^{-rt}, t \geq 0)$  is a martingale if and only if  $\mu + \lambda E(Y_1) = r$ .*

## 5.2 Martingales

We now denote by  $\nu$  the measure  $\nu(dy) = \lambda F(dy)$ , a  $(\lambda, F)$  compound Poisson process will be called a  $(\lambda, \nu)$  compound Poisson process.

**Proposition 5.3** *If  $X$  is a  $(\lambda, \nu)$  compound Poisson process, for any  $\alpha$  such that  $\int_{-\infty}^{\infty} e^{\alpha u} \nu(du) < \infty$ , the process*

$$Z_t^{(\alpha)} = \exp \left( \alpha X_t + t \left( \int_{-\infty}^{\infty} (1 - e^{\alpha u}) \nu(du) \right) \right)$$

*is a martingale and*

$$E(e^{\alpha X_t}) = \exp \left( -t \left( \int_{-\infty}^{\infty} (1 - e^{\alpha u}) \nu(du) \right) \right).$$

**Proposition 5.4** *Let  $X$  be a  $(\lambda, \nu)$  compound Poisson process, and  $f$  a bounded Borel function. Then, the process*

$$\exp \left( \sum_{k=1}^{N_t} f(Y_k) + t \int_{-\infty}^{\infty} (1 - e^{f(x)}) \nu(dx) \right)$$

*is a martingale. In particular*

$$E \left( \exp \left( \sum_{k=1}^{N_t} f(Y_k) \right) \right) = \exp \left( -t \int_{-\infty}^{\infty} (1 - e^{f(x)}) \nu(dx) \right)$$

For any bounded Borel function  $f$ , we denote by  $\nu(f) = \int_{-\infty}^{\infty} f(x)\nu(dx)$  the product  $\lambda E(f(Y_1))$ .

**Proposition 5.5** *Let  $X$  be a  $(\lambda, \nu)$  compound Poisson process. The process*

$$M_t^f = \sum_{s \leq t} f(\Delta X_s) \mathbb{1}_{\{\Delta X_s \neq 0\}} - t\nu(f)$$

*is a martingale. Conversely, suppose that  $X$  is a pure jump process and that there exists a finite positive measure  $\sigma$  such that*

$$\sum_{s \leq t} f(\Delta X_s) \mathbb{1}_{\{\Delta X_s \neq 0\}} - t\sigma(f)$$

*is a martingale for any  $f$ , then  $X$  is a  $(\sigma(1), \sigma)$  compound Poisson process.*

## 5.3 Hitting Times

Let  $X_t = bt + \sum_{k=1}^{N_t} Y_k$ .

Assume that the support of  $F$  is included in  $] - \infty, 0]$ .

The process  $(\exp(uX_t - t\psi(u)), t \geq 0)$  is a martingale, with

$$\psi(u) = bu - \lambda \int_{-\infty}^0 (1 - e^{uy}) F(dy).$$

Let  $T_x = \inf\{t : X_t > x\}$ . Since the process  $X$  has no positive jumps,  $X_{T_x} = x$ .

Hence  $E(e^{uX_{t \wedge T_x} - (t \wedge T_x)\psi(u)}) = 1$  and when  $t$  goes to infinity, one obtains

$$E(e^{ux - T_x\psi(u)} \mathbb{1}_{\{T_x < \infty\}}) = 1.$$

If  $\psi$  admits an inverse  $\psi^\sharp$ , one gets if  $T_x$  is finite

$$E(e^{-\lambda T_x}) = e^{-x\psi^\sharp(\lambda)}.$$

Setting  $Z_k = -Y_k$ , the random variables  $Z_k$  can be interpreted as losses for insurance companies. The process  $z + bt - \sum_{k=1}^{N_t} Z_k$  is called the Cramer-Lundberg risk process. The time  $\tau = \inf\{t : X_t \leq 0\}$  is the bankruptcy time for the company.

**One sided exponential law.** If  $F(dy) = \theta e^{\theta y} \mathbb{1}_{\{y < 0\}} dy$ , one obtains

$\psi(u) = bu - \frac{\lambda u}{\theta + u}$ , hence inverting  $\psi$ ,

$$E(e^{-\kappa T_x} \mathbb{1}_{\{T_x < \infty\}}) = e^{-x\psi^\#(\kappa)},$$

with

$$\psi^\#(\kappa) = \frac{\lambda + \kappa - \theta b + \sqrt{(\lambda + \kappa - \theta b)^2 + 4\theta b}}{2b}.$$

## 5.4 Change of Measure

Let  $X$  be a  $(\lambda, \nu)$  compound Poisson process,  $\tilde{\lambda} > 0$  and  $\tilde{F}$  a probability measure on  $\mathbb{R}$ , absolutely continuous w.r.t.  $F$  and  $\tilde{\nu}(dx) = \tilde{\lambda}\tilde{F}(dx)$ . Let

$$L_t = \exp \left( t(\lambda - \tilde{\lambda}) + \sum_{s \leq t} \ln \left( \frac{d\tilde{\nu}}{d\nu} \right) (\Delta X_s) \right).$$

Set  $dQ|_{\mathcal{F}_t} = L_t dP|_{\mathcal{F}_t}$ .

**Proposition 5.6** *Under  $Q$ , the process  $X$  is a  $(\tilde{\lambda}, \tilde{\nu})$  compound Poisson process.*

## 6 An Elementary Model of Prices including Jumps

Suppose that  $S$  is a stochastic process with dynamics given by

$$dS_t = S_{t-}(b(t)dt + \phi(t)dM_t), \quad (6.1)$$

where  $M$  is the compensated compensated martingale associated with a Poisson process and where  $b, \phi$  are deterministic continuous functions, such that  $\phi > -1$ . The solution of (6.1) is

$$\begin{aligned} S_t &= S_0 \exp \left[ - \int_0^t \phi(s)\lambda(s)ds + \int_0^t b(s)ds \right] \prod_{s \leq t} (1 + \phi(s)\Delta N_s) \\ &= S_0 \exp \left[ \int_0^t b(s)ds \right] \exp \left[ \int_0^t \ln(1 + \phi(s))dN_s - \int_0^t \phi(s)\lambda(s)ds \right]. \end{aligned}$$

Hence  $S_t \exp\left(-\int_0^t b(s)ds\right)$  is a strictly positive martingale.

We denote by  $r$  the deterministic interest rate and by  $R_t = \exp\left(-\int_0^t r(s)ds\right)$  the discounted factor.

Any strictly positive martingale  $L$  can be written as  $dL_t = L_{t-}\mu_t dM_t$  with  $\mu > -1$ .

Let  $(Y_t = R_t S_t L_t, t \geq 0)$ . Itô's calculus yields to

$$\begin{aligned} dY_t &\cong Y_{t-} ((b(t) - r(t))dt + \phi(t)\mu_t d[M]_t) \\ &\cong Y_{t-} (b(t) - r(t) + \phi(t)\mu_t \lambda(t)) dt. \end{aligned}$$

Hence,  $Y$  is a local martingale if and only if  $\mu_t \lambda(t) = -\frac{b(t) - r(t)}{\phi(t)}$ .

Assume that  $\mu > -1$  and  $Q|_{\mathcal{F}_t} = L_t P|_{\mathcal{F}_t}$ . Under  $Q$ ,  $N$  is a Poisson process with intensity  $((\mu(s) + 1)\lambda(s), s \geq 0)$  and

$$dS_t = S_{t-} (r(t)dt + \phi(t)dM_t^\mu)$$

where  $(M^\mu(t) = N_t - \int_0^t (\mu(s) + 1)\lambda(s) ds, t \geq 0)$  is the compensated  $Q$ -martingale.

Hence  $Q$  is the unique equivalent martingale measure.

## 6.1 Price Process

Let

$$dS_t = (\alpha S_{t-} + \beta) dt + (\gamma S_{t-} + \delta) dX_t \quad (6.2)$$

where  $X$  is a  $(\lambda, \nu)$  compound Poisson process. The solution of (6.2) is a Markov process with infinitesimal generator

$$\mathcal{L}(f)(x) = \int_{-\infty}^{+\infty} [f(x + \gamma xy + \delta y) - f(x)] \nu(dy) + (\alpha x + \beta) f'(x).$$

Let  $S$  be the solution of (6.2). The process  $e^{-rt} S_t$  is a martingale if and only if

$$\gamma \int y \nu(dy) + \alpha = r, \quad \delta \int y \nu(dy) + \beta = 0.$$

Let  $\tilde{F}$  be a probability measure absolutely continuous with respect to  $F$  and

$$L_t = \exp \left( t(\lambda - \tilde{\lambda}) + \sum_{s \leq t} \ln \left( \frac{\tilde{\lambda}}{\lambda} \frac{d\tilde{F}}{dF} \right) (\Delta X_s) \right).$$

Let  $dQ|_{\mathcal{F}_t} = L_t dP|_{\mathcal{F}_t}$ . The process  $(S_t e^{-rt}, t \geq 0)$  is a  $Q$ -martingale if and only if

$$\tilde{\lambda}\gamma \int y \tilde{F}(dy) + \alpha = r, \quad \tilde{\lambda}\delta \int y \tilde{F}(dy) + \beta = 0$$

Hence, there are an infinite number of e.m.m.. One can change the intensity of the Poisson process, and/or the law of the jumps.